# **Archival Report**

## A Longitudinal Imaging Genetics Study of Neuroanatomical Asymmetry in Alzheimer's Disease

Christian Wachinger, Kwangsik Nho, Andrew J. Saykin, Martin Reuter, and Anna Rieckmann for the Alzheimer's Disease Neuroimaging Initiative

### ABSTRACT

**BACKGROUND:** Contralateral brain structures represent a unique, within-patient reference element for disease, and asymmetries can provide a personalized measure of the accumulation of past disease processes. Neuroanatomical shape asymmetries have recently been associated with the progression of Alzheimer's disease (AD), but the biological basis of asymmetric brain changes in AD remains unknown.

**METHODS:** We investigated genetic influences on brain asymmetry by identifying associations between magnetic resonance imaging-derived measures of asymmetry and candidate single nucleotide polymorphisms (SNPs) that have previously been identified in genome-wide association studies for AD diagnosis and for brain subcortical volumes. For analyzing longitudinal neuroimaging data (1241 individuals, 6395 scans), we used a mixed effects model with interaction between genotype and diagnosis.

**RESULTS:** Significant associations between asymmetry of the amygdala, hippocampus, and putamen and SNPs in the genes *BIN1*, *CD2AP*, *ZCWPW1*, *ABCA7*, *TNKS*, and *DLG2* were found.

**CONCLUSIONS:** The associations between SNPs in the genes *TNKS* and *DLG2* and AD-related increases in shape asymmetry are of particular interest; these SNPs have previously been associated with subcortical volumes of amygdala and putamen but have not yet been associated with AD pathology. For AD candidate SNPs, we extend previous work to show that their effects on subcortical brain structures are asymmetric. This provides novel evidence about the biological underpinnings of brain asymmetry as a disease marker.

Keywords: Alzheimer's, Asymmetry, Genetics, Imaging, Longitudinal, Shape

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Alzheimer's disease (AD) is a progressive and irreversible brain disorder that is characterized by a gradual degradation of cognitive functions over many years that includes a long preclinical phase (1). Because brain atrophy as measured with magnetic resonance imaging (MRI) correlates with neuron loss (2), longitudinal in vivo neuroimaging has become invaluable for studying trajectories of pathophysiological change in AD. Repeated volumetric measurements of brain volumes in the same individuals have shown accelerated rates of atrophy in patients with AD compared with healthy control subjects, even in preclinical stages (3).

Volume measurements are, however, only a crude simplification of the complex anatomical change that occurs in aging and AD and often ignore the fact that atrophy is not uniform across a brain structure, e.g., the hippocampus (4–6). In contrast, shape descriptors are sensitive to such changes as they retain more geometrical information (7). Indeed, a recent study revealed that subtle preclinical changes in the shape asymmetry of subcortical brain structures could predict the conversion from mild cognitive impairment (MCI) to dementia more accurately than volumetric asymmetry (8). These asymmetries are undirectional, i.e., they do not have a consistent hemispheric effect and therefore refer to the magnitude of asymmetry independent of direction.

While structural asymmetries could serve as imaging biomarkers for the early presymptomatic classification and prediction of AD, possible biological mechanisms that underlie asymmetric manifestation of AD pathology are unclear. Here, we investigate the genetic influence on shape asymmetry in AD in an imaging quantitative trait loci (QTL) analysis including healthy control subjects and MCI stable, MCI progressor, and AD patients. While previous genome-wide association studies (GWASs) have revealed single nucleotide polymorphisms (SNPs) that are related to AD diagnosis, the mechanism through which they affect the disease remains largely unknown. Relating these same SNPs to imaging markers helps our understanding of how common genetic variants alter specific structures and pathways in the living human brain (9). Because previous research is inconsistent with respect to heritability of brain structural asymmetry (10-12), we model

both main effects of SNP as well as interactions between diagnosis groups and SNP on brain asymmetry. Whereas main effects speak to heritability per se, a significant interaction would reveal genetic influences on brain shape asymmetry that are magnified in AD patients and therefore can be interpreted to reflect the development of disease (13).

Imaging QTL studies may have several potential advantages over case-control studies, including increased power (14). Imaging endophenotypes of disease in QTL studies can separate patients and normal subjects more accurately and therefore limit the confound of including asymptomatic subjects in the control group, which is particularly important for the long clinically silent prodromal phase of AD (13,15–21). These previous imaging QTL studies on AD risk variants and MRI measures used a cross-sectional design and focused on volume of brain structures and cortical thickness. We use a longitudinal model to study the genetics of shape asymmetry, yielding increased power for detecting genetic associations (22), although this may also depend on the genetic architecture of the specific traits being studied.

The current study uses longitudinal imaging data from more than 6000 MRI scans and genetic data from 1241 individuals in the Alzheimer's Disease Neuroimaging Initiative (ADNI). We focus on four brain structures (the hippocampus, amygdala, putamen, and caudate) that were selected a priori based on previous reports of increased volume and shape asymmetries in AD (8,23). Shape asymmetry is computed with the Mahalanobis distance of lateralized brain structures within a subject. We include 31 SNPs, selected a priori based on previous GWASs results. These SNPs are primarily composed of 21 candidate SNPs that have been associated with late-onset AD in GWASs (24-26). If these genetic risk variants for AD also influence shape asymmetry in AD, the results could suggest mechanisms or pathways through which these genes might be exerting their influence in AD. In a more exploratory analysis, 10 additional SNPs are included that have recently been associated to subcortical volume in a large-scale GWAS (27), but not to AD per se. The motivation is to identify possible genetic predispositions that render the brain vulnerable to shape asymmetry in disease. For example, genes that are associated with smaller hippocampi might render the hippocampus more vulnerable to AD pathology even if the genes are not directly implicated in AD pathology.

#### **METHODS AND MATERIALS**

### Data

We analyzed data from the ADNI, which was launched in 2003 as a public-private partnership, led by principal investigator Michael W. Weiner. The primary goal of ADNI has been to test whether serial MRI, positron emission tomography, other biological markers, and clinical and neuropsychological assessment can be combined to measure the progression of MCI and early AD (for up-to-date information, visit www.adni-info.org). We select all subjects with genetic information and at least three longitudinal MRI scans from the ADNI cohort, yielding 1241 individuals and 6395 scans with summary statistics listed in Supplemental Table S1. Supplemental Figure S1 shows a histogram of the number of scans per subjects.

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### **Analysis of Brain Structure Shape**

We computed the brain structure shape based on the brain descriptor BrainPrint, which relies on automated segmentation with FreeSurfer (28–31). BrainPrint and its application to study shape asymmetry has previously been described in detail (7,8). Briefly, after image segmentation, geometric representations are extracted for the identified subcortical structures via the marching cubes algorithm. shapeDNA (32) is used as the shape descriptor of the individual structures in the BrainPrint, which performs among the best in a comparison of methods for nonrigid three-dimensional shape retrieval (33). shapeDNA is based on the eigenvalues of the Laplace–Beltrami operator  $\varDelta$  can be computed via finite element analysis by solving the Laplacian eigenvalue problem (Helmholtz equation) on the given shape

$$\Delta f = -\lambda f.$$

The solution consists of eigenvalue  $\lambda_i \in \mathbb{R}$  and eigenfunction  $f_i$  pairs. The first *I* nonzero eigenvalues form the descriptor:  $\overline{\lambda} = (\lambda_1, ..., \lambda_l)$ , where we set l = 50 (7). A key property of the eigenvalues is their isometry invariance, i.e., length-preserving deformations will not change the spectrum. Isometry invariance includes rigid body motion as well as reflections and therefore permits direct comparison of shapes across individuals without any registration. The collection of shape descriptors from cortical and subcortical structures forms the BrainPrint, which has recently shown high potential for the automated diagnosis of dementia (34,35).

#### **Brain Asymmetry from BrainPrint**

Asymmetry of a lateralized brain structure s is measured by directly computing the Mahalanobis distance between the descriptors

$$Y_s = \| \ \overline{\lambda}_s^{\ \text{left}} {-} \overline{\lambda}_s^{\ \text{right}} \|_{\Sigma},$$

where we use a diagonal covariance matrix  $\Sigma$  with the *i*-th element  $\Sigma_{ii} = i^2$  to reduce the impact of higher eigenvalues on the distance (7). The asymmetry computation completely avoids lateral processing bias as it works on both hemispheres independently. The asymmetry measure presents a withinsubject measure that can identify directional and undirectional asymmetry. In addition, it allows for quantifying localized asymmetries potentially induced by morphometric changes in subnuclei, which can be crucial for tracking the progression of dementia (4,5).

### **Genetic Data**

In ADNI, genotyping was performed using three different Illumina platforms [Illumina Human610-Quad, Illumina Human-Omni Express, and Ilumina Omni2.5M BeadChips (36), Illumina, Inc., San Diego, CA]. The *APOE*  $\varepsilon$ 4 allele–defining SNPs rs429358 and rs7412 were separately obtained using standard methods (36). Standard quality control procedures for genetic markers and subjects are performed as described previously: 1) for SNP, SNP call rate < 95%, Hardy-Weinberg equilibrium test  $p < 1 \times 10^{-6}$ , and minor allele frequency

< 1%; and 2) for subject, subject sex, identity check, and subject call rate <95% (37). Because of the impact of population stratification on association analysis, we selected only non-Hispanic Caucasian participants using multidimensional scale analysis and HapMap GWAS genotypes (38). As the ADNI used different genotyping platforms, we imputed ungenotyped SNPs separately in each platform using MaCH (39) with the reference panel of the Haplotype Reference Consortium. After the imputed genotypes. From the imputed data, we selected 21 candidate AD SNPs (25) and 10 SNPs that have been associated with subcortical brain structures (27) based on a cut-off of  $p < 1 \times 10^{-7}$  (Supplement).

### **Statistical Analysis**

We used linear mixed effects models (40,41) to study the association between longitudinal change in brain asymmetry (i.e., the lateral shape distance) and genetics. Genotypes were coded as 0, 1, or 2, representing the number of minor alleles in the genotype, following an additive genetic model. We denoted age at baseline for individual *i* with  $B_i$ , years-from-baseline at follow-up scan *j* with  $X_{ij}$ , diagnosis with  $D_i$ , and the additive encoding of the SNP with  $S_i$ . To establish whether an association between SNP and asymmetry differs between diagnosis groups, the effect of interest in the model is the interaction term SNP × diagnosis. The linear model for asymmetry  $Y_{ij}$  as dependent variable is

$$Y_{ij} = \beta_0 + \beta_1 B_i + \beta_2 X_{ij} + \beta_3 S_i + \beta_4 D_i + \beta_5 S_i D_i + b_{0i} + b_{1i} X_{ij},$$

where  $\beta_0,..., \beta_5$  are fixed effects regression coefficients and  $b_{0i}$ ,  $b_{1i}$  are random effects regression coefficients. The random effects enable modeling individual-specific intercept and slope with respect to the time from the baseline. The fixed effect coefficient  $\beta_2$  models the longitudinal change on a population level. The statistical model is an adaptation of previous longitudinal models on the ADNI (8,40). We further evaluate a simplified model without an interaction between SNP and diagnosis. The following additional parameters are included as fixed effects (not shown in the equation above): years of education, sex, intracranial volume, and the number of *APOE*  $\varepsilon$ 4 risk alleles.

For the diagnosis variable, we differentiate between control subjects, MCI subjects that remain stable, MCI subjects that progress to AD, and AD patients. In our analysis, we encoded the diagnosis once as a quantitative variable and once as a categorical variable with four levels. The continuous coding of diagnosis has been used by Furney et al. (17) and Khondoker et al. (42), who also investigated the interaction SNP  $\times$  diagnosis. In their cross-sectional analyses, they have not differentiated between MCI progressor and stable. In a post hoc analysis, we include the interaction SNP imes years-frombaseline to the model to evaluate whether SNPs might have time-varied effects on asymmetry. We use false discovery rate (FDR) with q = .05 (43) to control for multiple hypothesis testing across SNPs similar to Li et al. (19), but also discuss Bonferroni correction in the Results section. Details on the implementation of the models appear in the Supplement.

### RESULTS

In the following, we describe the results of the SNPasymmetry analyses, where we used models with and without the interaction of SNP and diagnosis together with a quantitative and categorical coding of the diagnosis. Table 1 summarizes all the SNPs that showed significant associations in the different models together with their closest genes, location, major/minor alleles, minor allele frequency, genotype count, population-attributable fractions, or preventive fractions.

Significant interactions are reported in more detail, grouped by whether they were identified with a quantitative coding of disease (0 = control [CN]), 1 = MCI stable, 2 = MCI progressor, 3 = AD) (Table 2) or categorical coding (Table 3). We show standardized regression coefficients and *p* values for the main effects and, in addition, adjusted *p* values after FDR correction for the interaction. Results are only included in Tables 2 and 3 when the adjusted *p* value of the interaction is < .05.

Interactions with a quantitative coding of disease reveal genetic variants that are associated with shape asymmetry in a stage-dependent manner. Significant associations exist between amygdala asymmetry and rs117253277, as well as between putamen asymmetry and rs683250 and rs6733839 (Table 2). The main SNP effect is not significant for any of these associations after FDR correction. The main diagnosis effect is highly significant for rs117253277 and rs6733839, but not for rs683250. Notably, all regression coefficients for diagnosis are positive, which indicates an increase in asymmetry with the progression of dementia, consistent with our previous results (8). rs117253277 shows a negative coefficient for SNP (-0.539), which means that the presence of a minor allele A decreases the asymmetry. Importantly, the positive interaction (coefficient estimate = 0.585) signifies that asymmetry increases with the number of minor alleles for demented subjects. For rs683250, the pattern is inverted, with minor alleles yielding an increase in asymmetry in controls but a decrease in the demented population. The SNPs (rs117253277 and rs683250) were identified in the subcortical GWASs for amygdala and putamen, respectively, which is consistent with the structures in which we observe diseasedependent associations with asymmetry; rs6733839 was identified in the AD GWASs.

Table 3 reports the results for the categorical coding of diagnosis. The categorical coding is less hypothesis driven than the continuous coding because it allows for nonlinear interactions that are driven by only two groups (CN  $\rightarrow$  MCI stable,  $CN \rightarrow MCI$  progressor, and  $CN \rightarrow AD$ ). That said, this analysis only revealed significant coefficients for the factor  $CN \rightarrow AD$ . Factors that distinguish MCI groups from CN do not show significant results for the interaction. The main SNP effect is not significant for any of these associations, where the main diagnosis effect is highly significant for all. As for the continuous model, the interactions of diagnosis with rs117253277 and rs6733839 are significant. In addition,  $SNP \times diagnosis$  interactions are obtained for hippocampal asymmetry for rs1476679 and rs4147929. Both SNPs have been reported in AD GWASs. Interestingly, they have an inverted effect on asymmetry, with a positive interaction

Table 1. Su	Table 1. Summary of SNPs With Significant Associations With Brain Asymmetry	s With Sign	ificant Associá	ations With Bra	ain Asymr	netry					
				Major/							
SNP	Chromosome Position Closest G	Position	Closest Gene	Minor Alleles GWAS <sup>a</sup> Asymmetry <sup>b</sup>	GWAS <sup>a</sup>	Asymmetry <sup>b</sup>	Effect <sup>c</sup>	PAF Type <sup>d</sup>	PAF (%)	PAF Type <sup>d</sup> PAF (%) Genotype Count <sup>e</sup>	MAF
rs6733839	0	127892810	BIN1	C/T	AD	Putamen	SNP  imes disease	Risk	8.1	2382/2977/1036	T = 0.3950
rs10948363	9	47487762	CD2AP	AG	AD	Amygdala	SNP	Risk	2.3	3375/2549/471	G = 0.1881
rs4147929	19	1063443	ABCA7	G/A	AD	Hippocampus	Hippocampus SNP $ imes$ disease	Risk	2.8	3337/2503/555	C = 0.2115
rs1476679	7	100004446	ZCWPW1	T/C	AD	Hippocampus	Hippocampus SNP × disease	Preventive	3.2	3.2 4369/1828/198	A = 0.1751

(25)] and a subcortical brain GWAS [Hibar et al.  $SNP \times disease$ SNP, Putamen Putamen from an AD GWAS [Lambert et al. G/A DLG2 SNPs included in the study were selected 83276168 ÷ s683250

A = 0.0264 T = 0.3950

2173/3318/904

6320/75/0

ı ı

I

 $SNP \times disease$ 

Amygdala

Amygdala

C/A

TNKS

9304257

ω

rs117253277

AD, Alzheimer's disease; GWAS, genome-wide association study; MAF, minor allele frequency; PAF, population-attributable fraction; SNP, single nucleotide polymorphism. (27)

<sup>a</sup>GWAS identification of SNPs. Brain structures for subcortical GWAS.

(25)

<sup>5</sup>Association of SNP to brain asymmetry identified in the current study

<sup>c</sup>Significant main or interaction effect in our model

<sup>2</sup>PAF type for AD risk SNPs [from Supplemental Table 6 in Lambert et al.

Genotype count for (0/1/2) for image scans. MAF based on the 1000 Genomes Project. "Genotype count

coefficient for rs1476679 (estimate = 0.255) and a negative coefficient for rs4147929 (estimate = -0.309). This is consistent with their respective role in AD, as reported in Table 1; rs4147929 is a risk locus, whereas rs1476679 is a preventive locus (25). The minor allele frequency and the genotype count of rs117253277 are low, which may bias the results. For confirmation, we created random samples of similar sample size that matched the diagnostic distribution. The estimates for the interaction SNP  $\times$  diagnosis over 50 repetitions are plotted in Supplemental Figure S2. The median of 2.08 and the mean of 2.01 are close to the estimate of the original model (2.36).

Figure 1 displays the estimated intra- and interindividual change of the lateral shape asymmetry for the hippocampus, amygdala, and putamen with the associated loci. We show the genotype for control subjects and AD patients. Solid lines depict the global age effect, where the offset in intercept is determined by the genotype. Short line ticks depict the longitudinal intraindividual effect. The common pattern, except for rs117253277, is that genotype has a limited effect on the asymmetry of control subjects but a strong effect for AD patients. For rs117253277, the number of minor alleles also influences the asymmetry of control subjects, which illustrates the strong main effect of SNP (-0.268) in Table 3. For the hippocampus and amygdala, we observe a higher intraindividual increase in asymmetry compared with the interindividual increase (i.e., the age effect), as previously reported (8). Note that cross-sectional and longitudinal effects can vary substantially in Figure 1, which may result from positive selection bias for very old adults in crosssectional studies.

Table 4 provides statistical detail for the model with main effect of SNP only for quantitative coding (see Supplemental Table S2 for categorical coding). The association of rs683250 to putamen asymmetry is consistent with the interaction. A new association with amygdala asymmetry is found for the AD candidate SNP rs10948363.

In a post hoc analysis, we added the interaction SNP imesyears-from-baseline to the models and evaluated whether the interaction was significant for the above identified pairings of asymmetry and SNP. In the model without SNP  $\times$  diagnosis interaction, we found that the interaction SNP imes years-frombaseline was significant for rs683250 and putamen asymmetry ( $\beta$  = .043, p = .00127). Figure 2 illustrates the intra- and interindividual change by genotype, which shows that minor alleles were associated with a steeper increase in asymmetry over time.

In all the presented analyses, we included the number of APOE4 risk alleles as a covariate. In an additional analysis, we removed it from covariates and consider it as the SNP of interest. Across all the models, with and without interaction, as well as the different coding of diagnosis, there were no significant associations between asymmetry and APOE4.

We used FDR correction to control for multiple testing, but almost all results would also be significant with the conservative Bonferroni correction (p value threshold of .00161). The only exception is the interaction of rs6733839 with diagnosis for putamen asymmetry in the quantitative coding, although the interaction would still be significant for the categorical

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Mixed Effects Model With Interactions												
			$\beta_3$ (SNP) $\beta_4$ (Diagnosis) $\beta_5$ (SNP $ imes$ Diagnosis)				Diagnosis)					
Asymmetry Area	SNP	GWAS	β	p Value	β	p Value	β	p Value	Adjusted p Value			
Amygdala	rs117253277	Amygdala	539	.04136	.225	.00000	.585	.00129	.03738			
Putamen	rs683250	Putamen	015	.76344	.026	.43458	.097	.00124	.03833			
Putamen	rs6733839	AD	.060	.18796	.180	.00000	086	.00314	.04864			

 Table 2. Standardized Regression Coefficients and p Values for the Analysis of Asymmetry With Genetic Loci for the Linear

 Mixed Effects Model With Interactions

Adjusted *p* values for the interaction are presented, where we only show significant associations after false discovery rate correction. *p* values are rounded to five decimal places. Diagnosis is modeled as a continuous variable.

AD, Alzheimer's disease; GWAS, genome-wide association study; SNP, single nucleotide polymorphism.

coding. Note that we do not correct for multiple comparisons across models.

### DISCUSSION

In a series of linear mixed effects models of longitudinal neuroanatomical change, we have identified genetic risk variants associated with an increase in brain shape asymmetry in AD. The closest genes associated with significant SNPs include BIN1 (rs6733839), CD2AP (rs10948363), and ABCA7 (rs4147929), which code for proteins involved in amyloid generation, secretion, and clearance and are thus likely directly implicated in the accumulation of AD pathology. Another gene. ZCWPW1 (rs1476679), has previously been identified to have a preventative effect in AD. Interestingly, we also identified SNPs in the genes TNKS and DLG2 as risk variants for AD-related increases in shape asymmetry. These are SNPs that have previously been associated with subcortical volumes of the amygdala and putamen, respectively (27), and we show that these same SNPs also convey risk for AD pathology in these same structures. Previous reports have found several associations between neuroimaging measures and APOE4 (19). In contrast, we found no significant associations between neuroanatomical asymmetry and APOE4, suggesting that the association between APOE4 and atrophy is global, i.e., not symmetric. We review the specifics of each significant SNP below before discussing our results more generally.

The interaction of rs117253277 (*TNKS*) with diagnosis showed the most significant association in our study ( $p = 6 \times 10^{-6}$ ). The SNP was identified in the subcortical GWAS as a common variant associated with differences in amygdala volume. Our results show that the SNP also influences amygdala asymmetry in the context of AD pathology, which provides novel evidence that inherent differences in amygdala volumes

make the brain more vulnerable to AD-related patterns of atrophy (i.e., increases in shape asymmetry). Importantly, the interaction between SNP and disease was significant when coding disease both categorically and continuously, which shows that the SNP promotes increases in shape asymmetry already in the preclinical stages of AD in a dose-dependent response. Additional experiments were performed on matched samples to confirm the reliable estimate despite the low minor allele frequency of the SNP. The closest gene is *TNKS*, which catalyzes the adenosine diphosphate ribosylation of target proteins.

rs683250 is an intronic locus within *DLG2* and was identified in the subcortical GWAS. It seems to predispose the brain to undergo asymmetric shape atrophy in the progression to AD. For rs683250, however, we also identified a significant main effect of SNP on shape asymmetry, suggesting that the related gene has an effect on putamen asymmetry per se that is magnified in disease. Moreover, we found in the post hoc analysis that there was a significant interaction between SNP and years-from-baseline, indicating that the SNP affects change in putamen asymmetry over time. Genetic variants in *DLG2* affect learning and cognitive flexibility (44) and are associated with schizophrenia (45). The link to schizophrenia of the SNP is interesting, because abnormal asymmetries in subcortical structures have previously been reported for schizophrenia (46,47).

rs6733839 is within the *BIN1* gene and the most important genetic susceptibility locus after *APOE4* for individuals of European ancestry (48). As rs117253277, the SNP was significant for categorical and continuous coding. GWAS studies with MRI measures found association of *BIN1* with atrophy in the hippocampus (16), entorhinal and temporal pole cortex (15), and left parahippocampal and right inferior parietal cortex (19). Functions of the *BIN1* gene include the production and clearance of

 Table 3. Standardized Regression Coefficients and p Values for the Analysis of Lateral Asymmetry With Genetic Loci for the

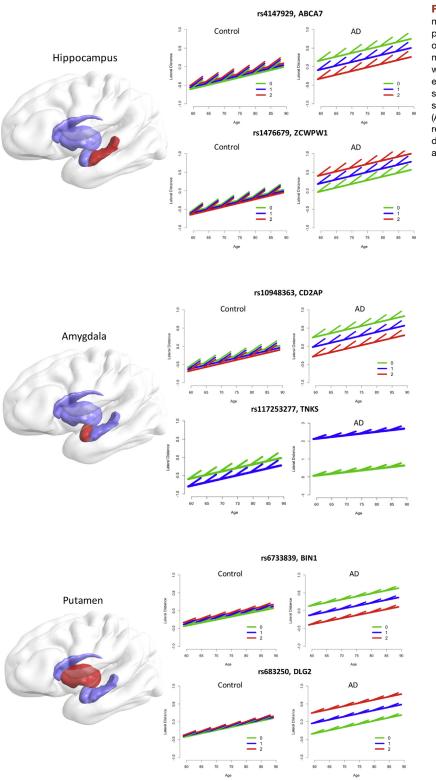
 Linear Mixed Effects Model With Interactions

			$\beta_3$	$\beta_3$ (SNP) $\beta_4$ (Diagnosis)			β <sub>5</sub> (	$eta_5$ (SNP $ imes$ Diagnosis) CN $ o$ AD			
Asymmetry Area	SNP	GWAS	β	p Value	β	<i>p</i> Value	β	<i>p</i> Value	Adjusted p Value		
Hippocampus	rs1476679	AD	037	.42071	.600	.00000	.255	.00148	.02213		
Hippocampus	rs4147929	AD	.060	.29610	.787	.00000	309	.00147	.02213		
Amygdala	rs117253277	Amygdala	268	.38490	.638	.00000	2.360	.00006	.00160		
Putamen	rs6733839	AD	.044	.41800	.607	.00000	323	.00051	.01363		

Adjusted *p* values for the interaction are presented, where we only show significant associations after false discovery rate correction. *p* values are rounded to five decimal places. Diagnosis is modeled as a categorical variable, where we only report the difference between CN and AD. The other factors were not significant.

AD, Alzheimer's disease; CN, control; GWAS, genome-wide association study; SNP, single nucleotide polymorphism.

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**Figure 1.** Longitudinal analysis of lateral asymmetry measures of the hippocampus, amygdala, and putamen with significantly associated single nucleotide polymorphisms. Lines and ticks illustrate estimates of the different linear mixed effects models with categorical coding of diagnosis. The global age effect is depicted by the slope of the long solid lines; short line ticks depict longitudinal slopes. Plots are shown for control subjects and Alzheimer's disease (AD) patients. The number of minor alleles can be related to higher or lower asymmetry in AD, depending on the single nucleotide polymorphism and its biological function.

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				$eta_3$ (SN	P)	β <sub>4</sub> (Ε	)iagnosis)					
Asymmetry Area	SNP	GWAS	β	p Value	Adjusted p Value	β	<i>p</i> Value					
Amygdala	rs10948363	AD	112	.00033	.01034	.235	.00000					
Putamen	rs683250	Putamen	.108	.00069	.02127	.107	.00000					

Table 4. Standardized Regression Coefficients and p Values for the Analysis of Lateral Asymmetry With Genetic Loci for the

Adjusted p values for the main effect SNP are presented, where we only show significant associations after false discovery rate correction. p values are rounded to five decimal places. Diagnosis is modeled as a continuous variable.

AD, Alzheimer's disease; GWAS, genome-wide association study; SNP, single nucleotide polymorphism.

amyloid- $\beta$  and cellular signaling; it increases the risk for AD by modulating tau pathology and is also involved in endocytosis, inflammation, calcium homeostasis, and apoptosis (49). Unlike the two AD-related SNPs related to hippocampus asymmetry, reviewed below, the SNP imes disease interaction for rs6733839 was significant for the continuous coding of disease groups, suggesting a linear increase in shape asymmetry with disease progression that includes preclinical stages.

For hippocampal asymmetry, two SNPs (rs1476679 and rs4147929) showed significant interaction with diagnosis. They have an inverse effect on asymmetry, which is consistent with their different roles as preventative or risk loci in AD. rs1476679 is intronic in the ZCWPW1 gene, is a histone modification reader, and is involved in genetic regulation (50,51). The preventative effect of rs1476679, reducing the risk of AD, was reported in Caucasians (25), a Spanish sample (52), and Han Chinese (53).

rs4147929 is within the ABCA7 gene. It is a transmembrane protein that influences neuronal cholesterol efflux and  $A\beta$ secretion (54). The gene is strongly expressed in hippocampus subfield CA1 (55) and associated with amyloid plaque burden (56). The minor allele of ABCA7 increases the risk of AD, as shown in an autopsy-confirmed research cohort (57). ABCA7 showed a significant association with hippocampal atrophy (20) and gray matter density (13). The expression of the gene in the subfield CA1 and the association with hippocampal atrophy are supportive of our results, because we have shown in previous work that the increase in asymmetry in hippocampus is not uniform but localized; this is one of the reasons for the improved results with shape descriptors compared with volumetric analyses.

There was only one SNP (rs10948363 [CD2AP]) that showed a significant association with shape asymmetry independent of disease status. This confirms previous reports that brain structural asymmetry per se is not strongly heritable (11),

but also shows that QTL studies can reveal subtle genetic influences that are not detectable in twin studies. CD2AP rs10948363 potentially contributes to amyloid precursor protein metabolism and subsequent amyloid- $\beta$  generation through its regulation of clathrin-mediated endocytosis (58). It is probably linked to modulating amyloid- $\beta$  clearance and tau neurotoxicity (56) and was associated with fluorodeoxyglucose positron emission tomography metabolism (13). Our results suggest its influence on amygdala asymmetry, which showed one of the strongest associations to dementia in our previous work (8).

Supplementary data from the UK Brain Expression Consortium on gene expression QTLs from postmortem healthy human brains (http://www.braineac.org/) revealed that homozygotes of the minor allele type for rs683250 show greater expression of DLG2 in the putamen of healthy individuals as compared with individuals carrying at least one major allele (Supplemental Figure S3). The results of the cis-eQTL mapping analysis are consistent with our results of the main effect of this SNP on putamen asymmetry.

Collectively, our results provide novel evidence for genes that may drive asymmetric accumulation of AD pathology and suggest that sequence variants may act through their influence on neuroanatomical asymmetry. It is important to note that brain structural asymmetry in AD, or disease more generally, is different from the lateralization of language and motor function, where higher asymmetry tends to relate to higher functioning. Instead, higher asymmetry in AD is associated with the progression of preclinical and prodromal stages of disease and reflects an asymmetric effect of pathologic processes on brain morphology. The core genetic mechanisms of lateralized human brain development are unknown (59). Future work may be directed at studying the relation between functional asymmetry in development and asymmetric disease manifestation.

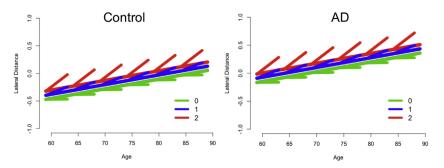


Figure 2. Longitudinal analysis for the single nucleotide polymorphism rs683250 and putamen asymmetry with interaction single nucleotide polymorphism  $\times$  years-from-baseline. Plots are shown for control and AD subjects for different genotype counts (0, 1, 2). A larger number of minor alleles corresponds with a steeper longitudinal increase. AD, Alzheimer's disease.

Several strengths and limitations of our work are worth noting. Working with the ADNI dataset is a strength because it is a large-scale, publicly available dataset that includes a rigorous clinical and genetic examination. However, ADNI was designed to simulate clinical trials and therefore uses more stringent inclusion and exclusion criteria, which necessitates replication on an independent sample in the general population. Another limitation of the study may be the a priori selection of candidate genes. A limitation of working with a global shape descriptor is that we cannot visualize the localized changes in shape asymmetry. A strength is that genetic associations of asymmetry are studied in a longitudinal design, where baseline asymmetry and change during the study period is modeled. This allowed us to not only model controls and AD patients, but further differentiate between MCI subjects that remain stable and those that progress to AD.

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Data used in preparation of this article were obtained from the ADNI database (adni.loni.usc.edu). As such, the investigators within the ADNI contributed to the design and implementation of ADNI and/or provided data but did not participate in analysis or writing of this report.

A complete listing of ADNI investigators can be found at http://adni.loni. usc.edu/wp-content/uploads/how\_to\_apply/ADNI\_Acknowledgement\_List. pdf.

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### **ARTICLE INFORMATION**

From the Laboratory for Artificial Intelligence in Medical Imaging (CW), Klinik für Kinder- und Jugendpsychiatrie, Klinikum der Universität München, Ludwig-Maximilians-Universität München, München, Germany; Center for Neuroimaging and Indiana Alzheimer Disease Center (KN, AJS), Department of Radiology and Imaging Sciences, Indiana University School of Medicine, Indianapolis, Indiana; A.A. Martinos Center for Biomedical Imaging (MR), Massachusetts General Hospital, Charlestown, Massachusetts; Deutsches Zentrum für Neurodegenerative Erkrankungen (MR), Bonn, Germany; and the Umeå Center for Functional Brain Imaging, Department of Radiation Sciences (AR), Umeå University, Umeå, Sweden.

Address correspondence to Christian Wachinger, Ph.D., Laboratory for Artificial Intelligence in Medical Imaging, Klinikum der Universität München, Ludwig-Maximilians-Universität München, Waltherstrasse 23, 80337 München, Germany; E-mail: christian.wachinger@med.uni-muenchen.de.

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